

Non-asymptotic error estimates for the Laplace approximation in Bayesian inverse problems

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Structure

Introduction

Central error estimate

Explicit error estimate

Perturbed linear problems with Gaussian prior

Focus

Approximation of posterior distribution of Bayesian inverse problem by **Gaussian distribution** according to Laplace's method.

Laplace's method

1. Replace log-posterior density by second order Taylor approximation around MAP estimate
2. Renormalise

Examples for use of Laplace approximation

- ▶ When **sampling** posterior distribution is **too expensive**.
- ▶ Inverse problems that are **close to linear problem**.

Motivation

- ▶ **Asymptotic properties** of Laplace approximation in small noise or large data limit have been studied extensively.
- ▶ In practice, one is often interested in **quantifying approximation error** for **given noise level**.

Problems

Nonlinearity of problem or **high problem dimension** may cause **large approximation error** even for low noise level.

Goal

Understand and **quantify influence of**

1. **nonlinearity of forward mapping,**
2. **problem dimension**

on Laplace approximation error.

Recent results

- ▶ **Asymptotic behaviour** of Laplace approximation **in context of inverse problems: Error in Hellinger distance** converges in order of noise level [Schillings, Sprungk, and Wacker 2020].
- ▶ **Bernstein–von Mises theorem** for inverse problems when **problem dimension tends to infinity** with certain rate as noise level tends to zero [Lu 2017].

Contribution

Main results

Non-asymptotic error estimates in total variation distance for Laplace approximation in Bayesian inverse problems:

1. Central error estimate
2. Error estimate that **makes explicit** influence of **non-Gaussianity of likelihood**, **non-Gaussianity of prior**, and **problem dimension**
3. Error estimate for **perturbed linear problems** with Gaussian prior that **makes explicit** influence of **nonlinear perturbation**

Total variation distance

Definition

Total variation distance between two probability measures μ and ν on $(\mathbb{R}^d, \mathcal{B}(\mathbb{R}^d))$ defined by

$$\begin{aligned}d_{\text{TV}}(\mu, \nu) &= \frac{1}{2} \int_{\mathbb{R}^d} \left| \frac{d\nu}{d\lambda} - \frac{d\mu}{d\lambda} \right| d\lambda \\ &= \sup_{A \in \mathcal{B}(\mathbb{R}^d)} |\nu(A) - \mu(A)|,\end{aligned}$$

where λ denotes Lebesgue measure on \mathbb{R}^d .

Total variation error of Laplace approximation is **measure of non-Gaussianity** of posterior distribution.

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Set-up

For $\varepsilon > 0$, recover $x \in \mathbb{R}^d$ from noisy measurement $y \in \mathbb{R}^d$, where

$$y = G(x) + \sqrt{\varepsilon}\eta.$$

- ▶ Nonlinear **forward mapping** G
- ▶ **Random noise** $\eta \in \mathbb{R}^d$ with standard normal distribution

$$\eta \sim \mathcal{N}(0, I_d)$$

- ▶ **Prior distribution**

$$\mu(dx) = \exp(-R(x))dx$$

- ▶ **Posterior distribution** given by Bayes' formula as

$$\mu^y(dx) \propto \exp\left(-\frac{1}{2\varepsilon}|y - G(x)|^2 - R(x)\right) dx$$

Laplace approximation

Assumption

$$I(x) := \frac{1}{2}|y - G(x)|^2 + \varepsilon R(x)$$

has **unique minimiser** $\hat{x} \in \mathbb{R}^d$, $I \in C^2(\mathbb{R}^d, \mathbb{R})$, and Hessian $(H)I(\hat{x})$ is **positive definite**.

Then, **Laplace approximation** of μ^y defined as

$$\mathcal{L}_{\mu^y} := \mathcal{N}(\hat{x}, \varepsilon \Sigma),$$

where $\Sigma := (H)I(\hat{x})^{-1}$. This way,

$$\begin{aligned}\mu^y(dx) &\propto \exp\left(-\frac{1}{\varepsilon}I(x)\right) dx, \\ \mathcal{L}_{\mu^y}(dx) &\propto \exp\left(-\frac{1}{2\varepsilon}\|x - \hat{x}\|_{\Sigma}^2\right) dx.\end{aligned}$$

Assumptions

Define $\Phi(x) := \frac{1}{2}|y - G(x)|^2$, so that $l(x) = \Phi(x) + \varepsilon R(x)$.

Bounds on log-likelihood and log-prior density

$\Phi, R \in C^3(\mathbb{R}^d, \mathbb{R})$ and there exists $K > 0$ such that

$$\max \left\{ \|D^3\Phi(x)\|_{\Sigma}, \|D^3R(x)\|_{\Sigma} \right\} \leq K \quad \text{for all } x \in \mathbb{R}^d,$$

where $\|D^3\Phi(x)\|_{\Sigma} := \sup \{ |D^3\Phi(x)(h_1, h_2, h_3)| : \|h_j\|_{\Sigma} \leq 1 \}$.

Quadratic bound on log-posterior density

There exists $0 < \delta \leq 1$ such that

$$l(x) - l(\hat{x}) \geq \frac{\delta}{2} \|x - \hat{x}\|_{\Sigma}^2 \quad \text{for all } x \in \mathbb{R}^d.$$

Want to estimate $d_{\text{TV}}(\mu^y, \mathcal{L}_{\mu^y})$ in terms of K, δ, d , and ε .

Central error estimate

Theorem

Under previous assumptions on Φ , R , and I , we have

$$d_{\text{TV}}(\mu^y, \mathcal{L}_{\mu^y}) \leq E_1(r_0; K) + E_2(r_0; \delta) \quad \text{for all } r_0 \geq 0,$$

where

$$E_1(r_0; K) := (2\varepsilon)^{-\frac{d}{2}} \frac{2}{\Gamma\left(\frac{d}{2}\right)} \int_0^{r_0} f(r) r^{d-1} \exp\left(-\frac{1}{2\varepsilon} r^2\right) dx,$$

$$f(r) := \exp\left(\frac{(1+\varepsilon)K}{6\varepsilon} r^3\right) - 1,$$

and

$$E_2(r_0; \delta) := \delta^{-\frac{d}{2}} \frac{\Gamma\left(\frac{d}{2}, \frac{\delta}{2\varepsilon} r_0^2\right)}{\Gamma\left(\frac{d}{2}\right)}.$$

Optimal choice of r_0

Proposition

Optimal choice of r_0 in previous estimate is either 0 or satisfies

$$\exp\left(\frac{(1+\varepsilon)K}{6\varepsilon}r_0^3\right) - 1 - \exp\left(\frac{1-\delta}{2\varepsilon}r_0^2\right) = 0.$$

Second error term can be written as

$$E_2(r_0; \delta) = (2\varepsilon)^{-\frac{d}{2}} \frac{2}{\Gamma\left(\frac{d}{2}\right)} \int_{r_0}^{\infty} r^{d-1} \exp\left(-\frac{\delta}{2\varepsilon}r^2\right) dx.$$

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Explicit error estimate

Theorem

Suppose that previous assumptions on Φ , R , and I hold. If K , δ , ε , and d satisfy

$$\frac{2}{e\delta^{\frac{d}{2}+\frac{3}{2}}} \exp\left(-\frac{1}{8}\left(\frac{6\delta^{\frac{3}{2}}}{(1+\varepsilon)\varepsilon^{\frac{1}{2}}K}\right)^{\frac{2}{3}}\right) \leq \frac{(1+\varepsilon)\varepsilon^{\frac{1}{2}}Kd^{\frac{3}{2}}}{6\delta^{\frac{3}{2}}} \leq \frac{1}{8},$$

then

$$d_{\text{TV}}(\mu^y, \mathcal{L}_{\mu^y}) \leq C(1+\varepsilon)\varepsilon^{\frac{1}{2}}K\Gamma_d,$$

where

$$C := \frac{2}{3}\sqrt{2}e \quad \text{and} \quad \Gamma_d := \frac{\Gamma\left(\frac{d}{2} + \frac{3}{2}\right)}{\Gamma\left(\frac{d}{2}\right)}.$$

► Note that $\Gamma_d \asymp \left(\frac{d}{2}\right)^{\frac{3}{2}}$ as $d \rightarrow \infty$.

Asymptotic behaviour as problem dimension $d \rightarrow \infty$

Index K_d , δ_d , and ε_d by $d \in \mathbb{N}$.

Corollary

Suppose that previous assumptions hold for all $d \in \mathbb{N}$. If $\delta_d \leq e^{-1/2}$, $\varepsilon_d \leq 1$,

$$\varepsilon_d^{\frac{1}{2}} K_d \rightarrow 0, \quad \text{and} \quad \varepsilon_d^{\frac{1}{2}} K_d d^{\frac{3}{2}} \leq 3 \left(\frac{\delta_d}{-8 \ln \delta_d} \right)^{\frac{3}{2}}$$

for all $d \in \mathbb{N}$, then for every $C > \frac{2}{3}\sqrt{2}e$ there exists $N \in \mathbb{N}$ such that

$$d_{\text{TV}}(\mu^y, \mathcal{L}_{\mu^y}) \leq C \varepsilon_d^{\frac{1}{2}} K_d d^{\frac{3}{2}}$$

for all $d \geq N$.

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Perturbed linear problems with Gaussian prior

- ▶ **Forward mapping** G given by linear mapping with small nonlinear perturbation of **size** $\tau \geq 0$,

$$G_\tau(x) = Ax + \tau F(x),$$

where $A \in \mathbb{R}^{d \times d}$ and $F \in C^3(\mathbb{R}^d)$.

- ▶ **Gaussian prior distribution** $\mu = \mathcal{N}(m_0, \Sigma_0)$

Assumption

There exists $\tau_0 > 0$, such that for all $\tau \in [0, \tau_0]$,

$$I_\tau(x) = \frac{1}{2} |Ax + \tau F(x) - y_\tau|^2 + \frac{\varepsilon}{2} \|x - m_0\|_{\Sigma_0}^2$$

has **unique minimiser** \hat{x}_τ with $(H I_\tau)(\hat{x}_\tau) > 0$. Furthermore, y_τ , \hat{x}_τ , and $\Sigma_\tau := H I_\tau(\hat{x}_\tau)^{-1}$ **converge** as $\tau \rightarrow 0$ with $\lim_{\tau \rightarrow 0} \Sigma_\tau > 0$.

Assumptions

Let $B(M) \subset \mathbb{R}^d$ denote a closed Euclidean ball with radius M around the origin.

Bounds on nonlinear perturbation

There exist $C_0, \dots, C_3 > 0$ and $M > 0$ such that

$$\|D^j F(x)\|_{\Sigma_\tau} \leq C_j, \quad j = 0, \dots, 3,$$

for all $x \in \mathbb{R}^d$ and $\tau \in [0, \tau_0]$, and

$$D^3 F \equiv 0 \quad \text{on } \mathbb{R}^d \setminus B(M).$$

Error estimate for perturbed linear problems

Theorem

Under the previous assumptions, there exists $\tau_1 \in (0, \tau_0]$ such that

$$d_{\text{TV}}(\mu^{y_\tau}, \mathcal{L}_{\mu^{y_\tau}}) \leq C\Gamma_d(1 + \varepsilon)\varepsilon^{\frac{1}{2}} \left(V(\tau)\tau + \frac{W}{2}\tau^2 \right)$$

for all $\tau \in [0, \tau_1]$, where

$$C := \frac{2}{3}\sqrt{2}e, \quad \Gamma_d := \frac{\Gamma(\frac{d}{2} + \frac{3}{2})}{\Gamma(\frac{d}{2})},$$

$$V(\tau) := C_3(\|A\|M + |y_\tau|) + 3C_2 \left\| A\Sigma_\tau^{\frac{1}{2}} \right\|,$$

$$W := C_3C_0 + 3C_2C_1.$$

Moreover, $\{V(\tau)\}_{\tau \in [0, \tau_1]}$ is bounded.

Conclusion

Laplace approximation in Bayesian inverse problems

- ▶ Have non-asymptotic bound for total variation error.
- ▶ Under certain conditions, total variation error depends **linearly** on non-Gaussianity of likelihood and prior.
- ▶ For perturbed linear problems, total variation error depends **linearly** on size of nonlinear perturbation.

Outlook

- ▶ Estimate error in Wasserstein distance to achieve better or no dependence on problem dimension.

References



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