Bayesian hypothesis testing in statistical inverse problems

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## Structure

### Introduction

### Maximum a posteriori testing

Definition and evaluation Interpretation as regularized test Optimality

### Performance under spectral source condition

A priori and a posteriori choice of prior covariance Numerical results

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# Set-up

Consider statistical linear inverse problem

$$Y = T u^{\dagger} + \sigma Z,$$

where

- T: X → Y bounded linear forward operator between real separable Hilbert spaces X and Y,
- $u^{\dagger} \in \mathcal{X}$  unknown quantity of interest,
- $\sigma > 0$  noise level,
- $\blacktriangleright$  *Z* white Gaussian noise process on  $\mathcal{Y}$ .

For each  $g \in \mathcal{Y}$  one has access to real-valued Gaussian random variable

$$\langle Y,g\rangle = \left\langle Tu^{\dagger},g\right\rangle_{\mathcal{Y}} + \sigma \left\langle Z,g\right\rangle.$$

# Estimation and inference of features

- X, Y typically function spaces such as L<sup>p</sup>(Ω) or H<sup>s</sup>(Ω) on some domain Ω ⊆ ℝ<sup>d</sup>.
- Often one is not interested in whole function u<sup>†</sup> but in certain features of it such as modes, homogeneity, monotonicity, or support.
- Many features can be described by (family of) bounded linear functionals φ ∈ X\*.
- We perform inference for such features by means of statistical hypothesis testing. Specifically, we test

$$H_0: \left\langle arphi, u^{\dagger} 
ight
angle_{\mathcal{X}^* imes \mathcal{X}} \leq 0 \quad ext{against} \quad H_1: \left\langle arphi, u^{\dagger} 
ight
angle_{\mathcal{X}^* imes \mathcal{X}} > 0.$$

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Example 1: Support inference in deconvolution

T convolution operator

$$Tu = h * u$$

on  $L^2(\mathbb{R})$  with kernel *h*.

- **Question:** Is supp  $u^{\dagger} \cap (a, b) = \emptyset$ ?
- Under assumption that u<sup>†</sup> is nonnegative, φ := 1<sub>[a,b]</sub> describes feature of interest

$$\left\langle \varphi, u^{\dagger} \right\rangle_{L^2} = \int_a^b u^{\dagger}(x) \mathrm{d}x.$$

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# Example 2: Linearity inference

Direct noisy measurement

$$Y = f^{\dagger} + \sigma Z$$

of function  $f^{\dagger} \in H^1_0(0,1) \cap H^2(0,1)$ .

- Question: Is  $f^{\dagger}$  linear on  $(a, b) \subseteq (0, 1)$ ?
- For  $u \in L^2(0,1)$ , let Tu = f be weak solution to

$$-f'' = u$$
 on  $(0,1)$ ,  $f(0) = f(1) = 0$ .

Under assumption that f<sup>†</sup> is concave, φ := 1<sub>[a,b]</sub> describes feature of interest

$$\left\langle \varphi, u^{\dagger} \right\rangle_{L^2} = -\int_a^b (f^{\dagger})''(x) \mathrm{d}x.$$

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Statistical properties of hypothesis tests

- Hypothesis test Ψ(Y) takes only values 0 (accepts) and 1 (rejects).
- Probability that test correctly rejects hypothesis H<sub>0</sub> should be large, i.e.,

 $\mathbb{P}_{u^{\dagger}}\left[\Psi(Y)=1\right]$ 

for  $u^{\dagger} \in \mathcal{X}$  that satisfies  $H_1$  (power of test).

Control probability that test falsely rejects hypothesis via

$$\sup\left\{\mathbb{P}_{u^{\dagger}}\left[\Psi(Y)=1
ight]:u^{\dagger}\in\mathcal{X} ext{ satisfies }H_{0}
ight\}$$

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(level of significance of test).

# Unregularized hypothesis testing

• Assume that  $\varphi \in \operatorname{ran} T^*$  and choose  $\Phi_0 \in \mathcal{Y}$  such that

$$T^*\Phi_0=arphi.$$

• Then  $\langle Y, \Phi_0 \rangle$  is **natural estimator** for desired quantity

$$\langle \varphi, u^{\dagger} \rangle_{\mathcal{X}} = \langle T^* \Phi_0, u^{\dagger} \rangle_{\mathcal{X}} = \langle \Phi_0, T u^{\dagger} \rangle_{\mathcal{Y}}.$$

Define test

$$\Psi_0(Y) := \mathbf{1}_{\langle Y, \Phi_0 \rangle > c}$$

• Test  $\Psi_0$  has level  $\alpha \in (0,1)$  and power

$$\mathbb{P}_{u^{\dagger}}\left[ \Psi_{0}(Y) = 1 
ight] = Q\left( Q^{-1}(lpha) + rac{\langle arphi, u^{\dagger} 
angle}{\sigma \left\| \Phi_{0} 
ight\|} 
ight),$$

for choice  $c := \sigma \|\Phi_0\| Q^{-1}(1-\alpha)$ , where Q is cdf of  $\mathcal{N}(0,1)$ .

# Problems

**Unregularized test**  $\Psi_0$  has two inherent **limitations**.

- 1. If  $\varphi \notin \operatorname{ran} T^*$ , approach **not** applicable.
- 2. Probe element  $\Phi_0$  is solution to **ill-posed equation**  $T^*\Phi_0 = \varphi$ . For certain features, norm of  $\Phi_0$  is huge, and **power** of  $\Psi_0$  is **arbitrarily close to level**.

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# Solutions

Both of these limitations can be overcome by **regularized hypothesis tests** 

$$\Psi_{\Phi,c}(Y) := \mathbf{1}_{\langle Y, \Phi \rangle > c}, \quad \Phi \in \mathcal{Y}, c \in \mathbb{R}.$$

- 1. Maximize (empirical) power among class of regularized level  $\alpha$  tests [Kretschmann, Wachsmuth, Werner 2022].
- 2. Define tests using **Bayesian approach**: Reject based upon posterior probabilities.
- 3. Choose probe element  $\Phi$  as **Tikhonov regularized solution** to equation  $T^*\Phi_0 = \varphi$ .

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# Bayesian set-up

Consider problem from Bayesian perspective,

 $Y = TU + \sigma Z.$ 

Assign Gaussian prior distribution Π = N(m<sub>0</sub>, C<sub>0</sub>) to U,
 C<sub>0</sub> symmetric, positive definite, trace class,
 U and Z independent.

Conditional distribution of U, given Y = y, almost surely Gaussian  $\mathcal{N}(m, C)$  with

$$C = \sigma^2 C_0^{\frac{1}{2}} \left( C_0^{\frac{1}{2}} T^* T C_0^{\frac{1}{2}} + \sigma^2 \mathrm{Id} \right)^{-1} C_0^{\frac{1}{2}},$$
  
$$m = m_0 + C_0^{\frac{1}{2}} \left( C_0^{\frac{1}{2}} T^* T C_0^{\frac{1}{2}} + \sigma^2 \mathrm{Id} \right)^{-1} C_0^{\frac{1}{2}} T^* (y - T m_0).$$

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# Maximum a posteriori testing

For  $\varphi \in \mathcal{X}$ , define maximum a posteriori (MAP) test  $\Psi_{MAP}$  by  $\Psi_{MAP}(y) := \begin{cases} 1 & \text{if } \mathbb{P}[\langle \varphi, U \rangle > 0 | Y = y] > \mathbb{P}[\langle \varphi, U \rangle \le 0 | Y = y], \\ 0 & \text{otherwise,} \end{cases}$   $= \begin{cases} 1 & \text{if } \mathbb{P}[\langle \varphi, U \rangle > 0 | Y = y] > \frac{1}{2}, \\ 0 & \text{otherwise.} \end{cases}$ 

Hypothesis H<sub>0</sub> needs to have positive prior probability.

▶ Conditional distribution of  $\langle \varphi, U \rangle_{\mathcal{X}}$ , given Y = y, is

 $\mathcal{N}(\langle \varphi, m \rangle_{\mathcal{X}}, \langle \varphi, C \varphi \rangle_{\mathcal{X}}).$ 

# Evaluating MAP test

• Cdf 
$$F_{\varphi}$$
 of  $\langle \varphi, U \rangle_{\mathcal{X}}$ , given  $Y = y$ , is

$$F_{arphi}(t) = \mathbb{P}\left[\langle arphi, U 
angle \leq t | Y = y
ight] = Q\left(rac{t - \langle arphi, m 
angle}{\langle arphi, C arphi 
angle^{1/2}}
ight),$$

where Q is cdf of  $\mathcal{N}(0, 1)$ .

Hence

$$egin{aligned} \Psi_{\mathsf{MAP}}(y) &= 1 & \Leftrightarrow & \mathbb{P}\left[\langle arphi, U 
angle_{\mathcal{X}} > 0 | Y = y
ight] > rac{1}{2} \ & \Leftrightarrow & F_{arphi}(0) < rac{1}{2} & \Leftrightarrow & \langle arphi, m 
angle_{\mathcal{X}} > 0. \end{aligned}$$

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Connection with Tikhonov regularization

We have

$$\langle \varphi, m \rangle_{\mathcal{X}} = \langle y, \Phi_{\mathsf{MAP}} \rangle - \langle m_0, T^* \Phi_{\mathsf{MAP}} - \varphi \rangle_{\mathcal{X}},$$

where

$$\Phi_{\mathsf{MAP}} := TC_0^{\frac{1}{2}} \left( C_0^{\frac{1}{2}} T^* TC_0^{\frac{1}{2}} + \sigma^2 \mathsf{Id} \right)^{-1} C_0^{\frac{1}{2}} \varphi.$$

► If T is compact and  $C_0$  commutes with  $T^*T$ , then  $\Phi_{MAP}$  is **minimizer** of

$$\Phi \mapsto \left\| T^* \Phi - \varphi \right\|_{\mathcal{X}}^2 + \sigma^2 \left\| C_0^{-\frac{1}{2}} V^* \Phi \right\|_{\mathcal{X}}^2,$$

where V is a unitary operator such that T = V |T|.

## Interpretation as regularized test

### Theorem [Kretschmann, Wachsmuth, Werner 2022]

Under a priori assumptions on  $u^{\dagger}$ , for every  $\varphi \in \overline{\operatorname{ran} T^*}$ ,  $\Phi \in \mathcal{Y}$ , and  $\alpha \in (0, 1)$ , rejection threshold  $c = c(\varphi, \Phi, \alpha)$  can be chosen such that regularized test

$$\Psi_{\Phi,c}(Y) = \mathbf{1}_{\langle Y,\Phi\rangle > c}$$

has **level**  $\alpha$  for testing  $H_0$  against  $H_1$ .

MAP test  $\Psi_{MAP}$  corresponds to regularized test  $\Psi_{\Phi_{MAP},c_{MAP}}$  with  $c_{MAP} := \langle m_0, T^* \Phi_{MAP} - \varphi \rangle_{\mathcal{X}}$  and has level  $\alpha$  if prior mean  $m_0$  is chosen according to

$$\langle m_0, T^* \Phi_{\mathsf{MAP}} - \varphi \rangle_{\mathcal{X}} = c(\varphi, \Phi_{\mathsf{MAP}}, \alpha).$$

# Optimality

### Theorem [Kretschmann, Wachsmuth, Werner 2022]

For  $\varphi \in \overline{\operatorname{ran} T^*}$  and under a priori assumptions on  $u^{\dagger}$ , there exists **optimal probe element**  $\Phi^{\dagger} \in \mathcal{Y}$  that **maximizes power** among all regularized level  $\alpha$  tests.

### Theorem

If T is compact with singular system  $(\tau_k, e_k, f_k)_{k \in \mathbb{N}}$  and if

$$\langle arphi, e_k 
angle_{\mathcal{X}} = 0 \quad ext{for all } k \in \mathbb{N} ext{ with } \langle T^* \Phi^\dagger, e_k 
angle_{\mathcal{X}} = 0,$$

then prior covariance  $C_0$  can be chosen such that power of  $\Psi_{MAP}$  is arbitrarily close to power of optimal regularized test  $\Psi_{\Phi^{\dagger},c(\varphi,\Phi^{\dagger},\alpha)}$ .

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A priori assumptions on  $u^{\dagger}$ 

### Assumptions

- 1. Forward operator T is Hilbert–Schmidt and injective.
- 2. Spectral source condition

$$u^{\dagger} = (T^*T)^{\frac{\nu}{2}}w, \quad \|w\|_{\mathcal{X}} \leq \rho$$

for some  $w \in \mathcal{X}$  and  $\nu, \rho > 0$ .

3. Prior covariance operator

$$C_0 = \gamma^2 (T^*T)^{\mu}$$

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for some  $\gamma > 0$  and  $\mu \ge 1$ .

# A priori choice of prior covariance

### Theorem

If prior covariance is chosen as  $C_0 = \gamma_0^2 \sigma^2 (T^*T)^{\mu}$  with  $\gamma_0 > 0$  and if  $\mu > \frac{\nu}{2} - 1$ , then **power** of  $\Psi_{MAP}$  is at least

$$\mathbb{P}_{u^{\dagger}}\left[\Psi_{\mathsf{MAP}}(Y)=1\right] \geq Q\left(Q^{-1}(\alpha) + \frac{\frac{\langle \varphi, u^{\dagger} \rangle}{\|\varphi\|} - 2\rho\gamma_{0}^{-\frac{\nu}{\mu+1}}}{\sigma\gamma_{0}^{\frac{1}{\mu+1}}}\right)$$

• Nontrivial power if feature size is above threshold  $2\rho\gamma_0^{-\frac{\nu}{\mu+1}}$ .

• Choose  $\gamma_0$  to maximize estimate for specific feature size.

A posteriori choice of prior covariance

MAP test Ψ<sub>MAP</sub> has **power** 

$$\mathbb{P}_{u^{\dagger}}\left[\Psi_{\mathsf{MAP}}(Y)=1\right]=Q\left(Q^{-1}(\alpha)-\frac{J_{\mathcal{T}u^{\dagger}}(\Phi_{\mathsf{MAP}}(\mathcal{C}_{0}))}{\sigma}\right),$$

where  $J_{\mathcal{T}u^{\dagger}} \colon \mathcal{Y} \to \mathbb{R}$  [Kretschmann, Wachsmuth, Werner 2022].

- Functional J<sub>Tu<sup>†</sup></sub> unaccessible, use empirical functional J<sub>Y</sub> instead.
- Choose  $C_0 = \gamma^2 (T^*T)^{\mu}$  and  $\gamma > 0$  as **minimizer** of

$$\gamma \mapsto J_Y(\Phi_{\mathsf{MAP}}(\gamma(T^*T)^{\mu})) + \omega(\log \gamma)^2$$

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with  $\omega > 0$ .

Due to dependence of Φ<sub>MAP</sub> on Y via γ, it is no longer guaranteed that test has level α.

# Numerical results - Deconvolution



Figure: Exact power of unregularized test (----), oracle MAP test (-----), and empirical power and level of MAP test (------) for  $\nu = 1$ ,  $\mu = 2$ ,  $\alpha = 0.1$ ,  $\omega = 0.003$ , and M = 1000 samples.

## Numerical results – Antiderivative problem



Figure: Exact power of unregularized test (----), oracle MAP test (-----), and empirical power and level of MAP test (------) for  $\nu = 1$ ,  $\mu = 2$ ,  $\alpha = 0.1$ ,  $\omega = 0.01$ , and M = 1000 samples.

# Conclusion

- MAP test based upon Gaussian prior can be evaluated via Tikhonov–Phillips regularization.
- ► MAP test is defined for any feature described by bounded linear functional φ ∈ X\*.
- Regularizing effect allows feature testing in noise regimes where unregularized testing is unfeasible.

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### Outlook

- Other choices of prior distribution.
- Further study of Tikhonov regularized tests.

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